



Stable equilibrium distribution of a LV system: An Approximate Message Passing approach

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Model and general problem

The model

Abundances of n living species $u : \mathbb{R}_+ \mapsto \mathbb{R}_+^n$ follow the generalized Lotka-Volterra ODE

$$\dot{u}(t) = u(t) \odot (r + (Y - I) u(t))$$

- $Y \in \mathbb{R}^{n \times n}$: **interaction matrix**
- $r \in \mathbb{R}_+^n$: **intrinsic growth rates** vector

Problem

The LV ODE has a **globally stable equilibrium** $u_* = [u_{*,i}] \in \mathbb{R}_+^n$ if \exists a diagonal matrix $C > 0$ such that $CY + Y^T C < 2C^2$.

We assume that

- Y and r are **random**, $r \perp\!\!\!\perp Y$
- $n \rightarrow \infty$
- $\limsup_n \|(Y + Y^T)/2\| < 1$ a.s.

Problem

Asymptotics of the **equilibrium distribution**, i.e., the random probability measure

$$\mu^{u_*} = \frac{1}{n} \sum_{i \in [n]} \delta_{u_{*,i}}$$

Usually $u_* \in$ border of 1st quadrant. Percentage of surviving species ?

LCP characterization of equilibrium

When $\|(Y + Y^T)/2\| < 1$, matrix $I - Y$ is a P -matrix, i.e., all principal minors > 0 . Then

- The **Linear Complementarity Problem** (LCP): here, find a vector z such that

$$\begin{aligned}z &\succcurlyeq 0 \\z \odot (r + (Y - I)z) &= 0 \\r + (Y - I)z &\preccurlyeq 0\end{aligned}$$

has an unique solution for each $r \in \mathbb{R}^n$

- u_* is this solution

Our approach

Asymptotics of μ^{u_*} when $u_* = \text{LCP}(Y - I, r)$

Interaction matrix from Gaussian Orthogonal Ensemble (GOE)

LV model

$$\dot{u} = u \odot (r + (\beta G - I) u)$$

where

- G is GOE.
- $r \perp\!\!\!\perp G$, empirical measure $\mu^r \xrightarrow{\text{a.s.}} \bar{\mu}_r$ is the Wasserstein space $\mathcal{P}_2(\mathbb{R}_+)$
- We take $\beta < 1/2$. Indeed, $\|G\| \xrightarrow{\text{a.s.}}_n 2$ as a GOE matrix. Thus u_* exists and the LCP problem is well-defined since $\limsup_n \beta \|G\| < 1$ a.s.

Theorem [AHMN'23]

Let $R \sim \bar{\mu}_r$ and $Z \sim \mathcal{N}(0, 1)$ with $R \perp\!\!\!\perp Z$. For each $\beta < 1/\sqrt{2}$, the system

$$\begin{aligned}\beta &= \frac{\alpha}{1 + \gamma\alpha^2} \\ \sigma^2 &= \alpha^2 \mathbb{E}(\sigma Z + R)_+^2 \\ \gamma &= \mathbb{P}[\sigma Z + R > 0]\end{aligned}$$

admits an unique solution $(\alpha, \sigma, \gamma) \in (\sqrt{2}, \infty) \times (0, \infty) \times (0, 1)$.
Moreover, for $\beta < 1/2$ ($< 1/\sqrt{2}$ in physics literature)

$$\mu^{u_*} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} \mathcal{L}((1 + \alpha^2\gamma)(\sigma Z + R)_+) \quad \text{in } \mathcal{P}_2(\mathbb{R}_+)$$

Note: γ is (a lower bound on) the limit proportion of surviving species

Result obtained previously by [Bunin'17], [Galla'18], ...

Proof by Approximate Message Passing (AMP)

AMP principle

Iterative algorithm widely studied in statistical physics, coding and wireless communications, learning theory, ...

Basic algorithm: G is a $n \times n$ GOE matrix, and f_0, f_1, \dots is a sequence of $\mathbb{R}^2 \rightarrow \mathbb{R}$ Lipschitz functions

General algorithm form

$$x_{k+1} = \left[x_{k+1, i} \right]_{i \in [n]} = G f_k(x_k, a) + \text{"correction"}, \quad f_k(x_k, a) = \left[f_k(x_{k, i}, a_i) \right]_i$$

where a is a parameter vector, and $(x_0, a) \perp\!\!\!\perp G$

Thanks to the correction, we can identify the asymptotics of the **joint empirical measure**

$$\mu^{a, x_1, x_2, \dots, x_k} = \frac{1}{n} \sum_{i \in [n]} \delta_{a_i, x_{1, i}, x_{2, i}, \dots, x_{k, i}} \in \mathcal{P}(\mathbb{R}^{k+1})$$

when $n \rightarrow \infty$, for each fixed k

AMP algorithm and results

AMP algorithm:

$$x_{k+1} = Gf(x_k, a) - \langle \partial_x f'_k(x_k, a) \rangle f_{k-1}(x_{k-1}, a)$$

with $\langle x \rangle = \sum x_i/n$ and $\partial_x f_k(x_k, a) = \left[\frac{d}{dx} f_k(x_{k,i}, a_i) \right]_i$

Approximation of so-called **message passing algorithms** in statistical physics

Assuming

$$\mu^{a, x_0} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} \mathcal{L}(\bar{a}, \bar{x}) \quad \text{in (say) } \mathcal{P}_2(\mathbb{R}^2),$$

$$\mu^{a, x_1, x_2, \dots, x_k} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} \mathcal{L}(\bar{a}) \otimes \mathcal{N}(0, R_k) \quad \text{in } \mathcal{P}_2(\mathbb{R}^{k+1})$$

where the sequence of covariance matrices (R_k) is constructed recursively according to the **State Evolution (SE) equations**

(Bayati, Montanari, Bolthausen, ... \sim 2010)

SE equations

Let (Z_1, Z_2, \dots) with $\mathcal{L}((Z_1, \dots, Z_k)) = \mathcal{N}(0, R_k)$ and $(Z_1, Z_2, \dots) \perp\!\!\!\perp (\bar{a}, \bar{x})$.

Recursion:

$$R_1 = \mathbb{E}f_0(\bar{x}, \bar{a})^2$$

...

Knowing $R_k = \text{Cov}(Z_1, \dots, Z_k)$, we have

$$\mathbb{E}[Z_{k+1}Z_\ell] = \begin{cases} \mathbb{E}f_k(Z_k, \bar{a})f_{\ell-1}(Z_{\ell-1}, \bar{a}) & \text{if } \ell = 2, \dots, k+1 \\ \mathbb{E}f_k(Z_k, \bar{a})f_0(\bar{x}, \bar{a}) & \text{if } \ell = 1 \end{cases}$$

$$\Rightarrow R_{k+1}$$

μ^{u_*} asymptotics: proof with AMP

Set $\alpha > 0$ and $a \in \mathbb{R}_+^n$ to be specified later. Put $f_0(x, a) = f_1(x, a) = \dots = f(x, a) = \alpha(x + a)_+$.

AMP algorithm

$$x_{k+1} = \alpha G(x_k + a)_+ - \alpha^2 \gamma_k (x_{k-1} + a)_+ \quad \text{with} \\ \gamma_k = \langle \partial_x (x_k + a)_+ \rangle = \langle \mathbb{1}_{x_k + a > 0} \rangle$$

Key observation: Using an idea of [Donoho-Montanari'13], [Montanari-Richard'16], we can show that (Z_k, Z_{k+1}) in the SE equations become **more and more correlated**:

$$\text{Corr}(Z_k, Z_{k+1}) \xrightarrow[k \rightarrow \infty]{} 1,$$

which implies that

$$\lim_k \text{aslim}_n \frac{\langle x_{k+1}, x_k \rangle}{\|x_{k+1}\| \|x_k\|} = 1.$$

Proof with AMP (cont'd)

Thus,

$$x_k = \alpha G(x_k + a)_+ - \alpha^2 \gamma_k (x_k + a)_+ + \varepsilon_k$$

where $\lim_k \text{aslim}_n \|\varepsilon_k\|^2/n = 0$. We rewrite this

$$x_k + a = \alpha G(x_k + a)_+ - \alpha^2 \gamma_k (x_k + a)_+ + a + \varepsilon_k, \text{ or}$$

$$-(x_k + a)_- = \alpha G(x_k + a)_+ - (1 + \alpha^2 \gamma_k)(x_k + a)_+ + a + \varepsilon_k, \text{ or}$$

$$-\frac{(x_k + a)_-}{1 + \alpha^2 \gamma_k} = \left(\frac{\alpha}{1 + \alpha^2 \gamma_k} G - I \right) (x_k + a)_+ + \frac{a}{1 + \alpha^2 \gamma_k} + \varepsilon'_k$$

in other words,

$$(x_k + a)_+ = \text{LCP} \left(\frac{\alpha}{1 + \alpha^2 \gamma_k} G - I, \frac{a}{1 + \alpha^2 \gamma_k} + \varepsilon'_k \right).$$

Remember that

$$u_* = \text{LCP}(\beta G - I, r)$$

Identifying the two, using LCP perturbation results + SE equations
 \Rightarrow the theorem

A more involved interaction matrix model

New interaction matrix model

Non-necessarily Gaussian, centered, variance profile, pairwise correlations, and sparse

$$Y_{ij} = \sqrt{v_{ij}}X_{ij}$$

- $\mathbb{E}X_{ij} = 0$, $\mathbb{E}X_{ij}^2 = 1$ and $(\mathbb{E}|X_{ij}|^k)^{1/k} \leq Ck^{\eta/2}$
- Elements of $\{X_{ii}, (X_{ij}, X_{ji})_{i < j}\}$ are independent
- $\text{Corr}(X_{ij}, X_{ji}) = \tau_{ij} \in [-1, 1]$ (correlation profile).

$V = [v_{ij}] \succcurlyeq 0$ is the variance profile matrix. For $K_n \gtrsim (\log n)^{\eta \vee 1}$,

- $\#$ non-zero elements per row $\leq \text{Cst} \times K_n$
- $v_{ij} \leq \text{Cst}/K_n$
- All row sums $\geq \text{Cst}$

Ecological interpretation

- Couples of pairwise interactions are centered and independent
- **Pairwise correlations** specific to couples (i, j) , reminiscent of the well-known **elliptic model**:
 - $\tau_{ij} = 1$: often models competitive or mutualistic interactions
 - $\tau_{ij} = -1$: predator-prey
 - $\tau_{ij} = 0$: uncorrelated interactions
- Variance profile V : **inhomogeneous interaction strengths**
- **Sparsity**: every species interacts with a **small proportion**, K_n/n , of other species

Asymptotic behavior of μ^{u^*}

No assumed structure on the variance profile or the correlation profile matrices $V = [v_{ij}]$ and $T = [\tau_{ij}]$

Therefore, μ^{u^*} has **no reason to converge**

However, we can show that there exists a **deterministic sequence** (μ_n) of probability measures that approximates μ^{u^*} for large n

The parameters of a measure μ_n will be obtained through the solution of a large **system of equations**

With additional assumptions, this system can be reduced to two integral equations

Result: system of equations

Additional assumptions:

- r is deterministic
- $\limsup_n \|(Y + Y^T)/2\| < 1$ a.s. Conditions for this in terms of matrices V and T can be deduced from literature, e.g., [Bandeira-Van Handel'16]
- V and $Q = \left[\sqrt{v_{ij}v_{ji}\tau_{ij}} \right]_{i,j=1}^n$ have row sum norms $< 1/4$

System of equations

For each integer $n > 0$, let $Z \sim \mathcal{N}(0, I_n)$. The system of $2n$ equations in $(\rho, \zeta) \in \mathbb{R}_+^n \times [-1, 1]^n$

$$\begin{cases} \rho = V \operatorname{diag}(1 + \zeta)^2 \mathbb{E}(\sqrt{\rho} \odot Z + r)_+^2 \\ \zeta = \operatorname{diag}(1 + \zeta) Q \operatorname{diag}(1 + \zeta) \mathbb{P}[\sqrt{\rho} \odot Z + r \geq 0] \end{cases}$$

admits an unique solution.

Result: $\mu^{u^*} \sim$ mixture of truncated Gaussians

Theorem [GHN'24]

Let ξ be the Gaussian vector

$$\xi = \left[\xi_i \right]_{i \in [n]} = \text{diag}(1 + \zeta) (\sqrt{p} \odot Z + r)$$

and define the deterministic probability measure $\mu_n = \mathcal{L}((\xi_\theta)_+)$, where θ is a uniformly distributed random variable on $[n]$, independent of Z . Then,

$$\text{dist}_2(\mu^{u^*}, \mu_n) \xrightarrow[n \rightarrow \infty]{\mathcal{P}} 0$$

(distance in the space $\mathcal{P}_2(\mathbb{R})$)

Proof: a new AMP algorithm

New AMP for the generalized elliptic sparse model with a variance profile.
Might be helpful in contexts other than ecology, e.g., in learning theory

Measurement matrix: $W = S^{\odot 1/2} \odot X$, where

- X is a above (with a correlation profile)
- $S = [s_{ij}] \succcurlyeq 0$ has same properties as V above:
For $K_n \gtrsim (\log n)^{\eta \vee 1}$,
 - # non-zero elements per row $\leq \text{Cst} \times K_n$
 - $s_{ij} \leq \text{Cst}/K_n$
 - All row sums $\geq \text{Cst}$

f_0, f_1, \dots sequence of $\mathbb{R}^2 \rightarrow \mathbb{R}$ functions with sufficient regularity conditions

Deterministic parameter vector $a \in \mathbb{R}^n$ and initial value vector $x_0 \in \mathbb{R}^n$.

Algorithm:

$$x_{k+1} = \left[x_{k+1,i} \right] = Wf_k(x_k, a) - \text{diag}(B \partial f_k(x_k, a)) f_{k-1}(x_{k-1}, a)$$

$$B = \left[\sqrt{s_{ij}s_{ji}\tau_{ij}} \right] \text{ and } \partial f_k(x_k, a) = \left[\frac{d}{dx} f_k(x_{k,i}, a_i) \right]_{i \in [n]}.$$

We define a centered Gaussian family $(Z_{k,i})_{k \geq 1, i \in [n]}$ through appropriate DE equations (below)

Theorem

For each $k \geq 1$ and each continuous test function φ with quadratic growth at most,

$$\frac{1}{n} \sum_{i \in [n]} \varphi(a_i, x_{1,i}, \dots, x_{k,i}) - \mathbb{E} \varphi(a_i, Z_{1,i}, \dots, Z_{k,i}) \xrightarrow[n \rightarrow \infty]{\mathcal{P}} 0$$

(amounts to asymptotic behavior of μ^{a, x_1, \dots, x_k} in probability in the \mathcal{P}_2 space)

A **combinatorial proof** along the lines of [Bayati-Lelarge-Montanari'15] building on the message passing structure of the algorithm.

Thank you

Questions?

The DE equations for our model

Distribution of the centered Gaussian family $(Z_{k,i})_{k \geq 1, i \in [n]}$.

The n sequences $Z_i = (Z_{k,i})_{k \geq 1}$ are independent. The covariance matrices R_i^k of the vectors $\vec{Z}_i^k = [Z_{1,i}, \dots, Z_{k,i}]$ are defined recursively in k as follows.

$$R_i^1 = \sum_{\ell \in [n]} s_{i\ell} f_0(x_{0,i}, a_i)^2 \quad \text{for } i = 1, \dots, n$$

...

Given R_i^k for $i = 1, \dots, n$,

$$H_i^k = \mathbb{E} \begin{bmatrix} f_0(x_{0,i}, a_i) \\ f_1(Z_{1,i}, a_i) \\ \vdots \\ f_k(Z_{k,i}, a_i) \end{bmatrix} \begin{bmatrix} f_0(x_{0,i}, a_i) & f_1(Z_{1,i}, a_i) & \cdots & f_k(Z_{k,i}, a_i) \end{bmatrix}$$

$$R_i^{k+1} = \sum_{\ell=1}^n s_{i\ell} H_\ell^k \quad \text{for } i = 1, \dots, n$$